

Supplemental Materials

Investigating the Cognitive Structure of Stereotypes: Generic Beliefs About Groups Predict Social Judgments Better Than Statistical Beliefs

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STUDY 1

Method

Participants

Across all studies, we excluded participants who failed an attention-check item by giving a rating below 90% when estimating how many Caucasians (Study 1) or White people (Studies 2–4) are humans. In Study 1, 13 participants were excluded for failing this attention check (excluded participants' *average* = 43.7%; included participants' *average* = 99.9%).

Selection of Stereotypes for Study 1

The stereotypes in Study 1 were identified from prior research. We reviewed the first three pages of Google Scholar results when searching separately for the terms “stereotypes,” “stereotype endorsement,” and “stereotype generic beliefs” and identified 45 unique group/feature pairings. Sixty-five users on MTurk then rated on 0–100 sliding scales (0 = *No one believes this*; 100 = *Everyone believes this*) the extent to which people in society typically believe the corresponding generic statements (e.g., “African Americans are athletic”). Group/feature pairings were used in the present study if they met or exceeded the pre-determined threshold of the scale midpoint (see Table 1 of the main text for full list). We also selected the five lowest scoring items (e.g., “men are soft-spoken”, “politicians are trustworthy”) for inclusion as fillers to reduce acquiescence bias. Finally, we generated one attention-check item (“Caucasians are humans”; see above).

Additional Measure in Study 1: Social Desirability

After completing the ratings of statistical beliefs, generic beliefs, and social judgments, participants completed a measure of social desirability to ensure that any effects were not due to participants' hesitation in estimating the prevalence of stereotypical traits and/or in agreeing with the generic statements about these stereotypes. Participants completed an abbreviated 13-item measure of social desirability (e.g., "I have never intensely disliked anyone"; "I never hesitate to go out of my way to help someone in trouble"; Crowne & Marlowe, 1960; Reynolds, 1982). Scores were summed ($M = 6.94$, $SD = 3.10$). This measure was only administered in Study 1.

Pilot Study: Testing the Role of Distinctiveness

A pilot study assessed the role of statistical beliefs about *distinctiveness*—that is, the extent to which a trait (e.g., athleticism) is present in a group (e.g., African Americans) *relative to other groups* (e.g., Caucasians). We tested whether these statistical beliefs are more closely associated with social judgments than are generic beliefs and statistical beliefs about the absolute prevalence of a trait *within* a group.

Following an established practice in cognitive psychology (e.g., Cree, McNorgan, & McRae, 2006; Khemlani et al., 2012), we operationalized distinctiveness as *cue validity*—that is, the likelihood with which a stereotypical trait uniquely predicts membership in a social category (e.g., "given that Person Y is athletic, what is the likelihood they are African American?"). To ensure that participants understood how to make these judgments, we provided them with several examples and the opportunity to practice (with feedback).

We collected distinctiveness beliefs, generic beliefs, (absolute) statistical beliefs, and social judgments (namely, expectations) for 30 stereotypes, randomly assigning 80 MTurk users to provide ratings for one of these measures (as in Khemlani et al., 2012).

Distinctiveness did not predict unique variance in social judgments ($b = .02$, $t = 0.69$,

$p = .49$) beyond generic beliefs and (absolute) statistical beliefs (see Khemlani et al., 2012, for similar results with mostly non-social stimuli). Accordingly, to reduce the burden on participants in the within-person design used in our main studies, we did not collect ratings of distinctiveness.

Results

Supplementary Table 1 lists the means and standard deviations for each of the individual items. Aggregate summary statistics are presented in the main text (Table 2).

Supplementary Analyses by Type of Group

We explored whether generic and statistical beliefs predict social judgments differently depending on whether the beliefs are about race/ethnicity (12 items), gender (7 items), or professions (8 items; see Table 1 in the main text). Analyses revealed that people's generic beliefs predicted social judgments more strongly than their statistical beliefs *for each of these subsets* ($bs < -.208$, $ps < .001$). Thus, regardless of whether stereotypes were about race/ethnicity, gender, or profession, people's social judgments (i.e., their expectations about unfamiliar individuals' traits) were more strongly related to their generic beliefs than their statistical beliefs.

Maximum-Likelihood Estimation

Mplus requires Bayesian estimation for cross-classification of random effects (participants and items), which is the most appropriate way to model data in which participants rated the same set of items in three different forms (i.e., statistical beliefs, generic beliefs, social judgments). In addition, Bayesian estimation is more powerful than frequentist estimation because it more effectively handles multicollinearity and deviations from normality. Nonetheless, as an alternative to Bayesian estimation, we re-conducted the analyses in a two-level random-effects model (maximum likelihood, robust standard error estimation). Generic beliefs ($\beta = .513$, $SE = .045$, $p < .001$) predicted social judgments more

strongly ($b = -.172$, $SE = .075$, $p = .021$) than statistical beliefs ($\beta = .340$, $SE = .036$, $p < .001$). Thus, results were no different from those conducted using Bayesian estimation.

Moderating Effects

Social Desirability. Although our within-person design ruled out the possibility that the difference between generic and statistical beliefs was due to differences *between* participants, we nevertheless explored whether the relationship of these beliefs with social judgments was moderated by participants' tendency to provide socially desirable responses. Social desirability did not directly predict social judgments ($\beta = .017$ [$-.030$, $.064$], $p = .244$), nor did it moderate the relationship of statistical beliefs ($\beta = .029$ [$-.011$, $.068$], $p = .077$) or generic beliefs ($\beta = -.025$ [$-.065$, $.016$], $p = .118$) with social judgments.

Participant Age. Because working memory and fluid reasoning abilities typically decline with age (e.g., Salthouse, 2015), we also investigated whether older (vs. younger) participants rely more on generic beliefs and less on statistical beliefs in their social judgments. Participant age significantly moderated the link between generic beliefs ($\beta = .083$ [$.044$, $.123$], $p < .001$), but not statistical beliefs ($\beta = -.016$ [$-.052$, $.020$], $p = .193$), and social judgments. Generic beliefs were more predictive of social judgments for older (+1 *SD*) participants ($b = .641$ [$.583$, $.699$], $p < .001$) relative to younger (-1 *SD*) participants ($b = .474$ [$.420$, $.527$], $p < .001$). Although the present data cannot differentiate between ageing vs. cohort effects, these results are nevertheless broadly consistent with the argument that generic beliefs are central to stereotype structure because they are cognitively simple.

STUDY 2

Method

Participants

Participants received \$1.00 for this study, and the following studies, due to inclusion of additional measures. Nineteen participants were excluded for failing the attention check

(excluded participants' *average* prevalence estimate for “White people are humans” = 47.6%; included participants' *average* > 99.9%).

Stereotype Elicitation

Thirty-three participants on MTurk were instructed to “list as many stereotypes about people and groups of people that come to mind,” regardless of whether they personally believed them to be true. Two coders independently read the items and generated a summary stereotype for any stereotype that was listed by at least three participants (e.g., the stereotype “women are weak” summarized participants' stereotypes that women are “delicate,” “fragile,” and “weak”). Of all 581 stereotypes listed, 353 were listed by at least three participants. Disagreements were resolved via discussion, including a third party when needed. The final list of 30 items (see right side of Table 1 in the main text) comprised 18 items that were identical between coders, 8 items that addressed the same theme but with different phrasings (e.g., “Jewish people are rich” vs. “Jewish people are wealthy”; discrepancies resolved by discussion), and 4 items generated by only one of the two coders that were included after discussion.

In Study 2, we retained participants' terminology of “Black people” and “White people” in the final list of stereotypes rather than using “African Americans” and “Caucasians” (as in the psychological literature reviewed in Study 1). We chose this new language to reflect the language used by participants: No participants referred to “African Americans” and only one of the 33 referred to “Caucasians.”

As in Study 1, our final list of stereotypes included counter-stereotypic filler items (e.g., “men are soft-spoken”) and an attention-check item (“White people are humans”). Also as before, participants completed a 1-min distracter task between measures. Presentation order of the measures was randomized; within each measure, item order was randomized.

Supplementary Table 2 lists the means and standard deviations for each of the

individual items. Replicating effects using a different set of test stimuli was important because it would speak against the possibility that the results of Study 1 were due to idiosyncrasies in the particular list of stereotypes used (see Westfall, Judd, & Kenny, 2015).

Cognitive Style Moderator: The Cognitive Reflection Test

Participants completed the Cognitive Reflection Test (CRT), which consists of three items assessing analytic vs. intuitive thinking (Frederick, 2005). The original items from the CRT were modified slightly to avoid participants' pre-knowledge of the answers (Finucane & Gullion, 2010):

- (1) "If it takes 2 nurses 2 minutes to measure the blood pressure of 2 patients, how long would it take 200 nurses to measure the blood pressure of 200 patients?"
- (2) "Soup and salad cost \$5.50 in total. The soup costs a dollar more than the salad. How much does the salad cost?"
- (3) "Sally is doing an experiment on fruit fly reproduction. Every day, her fruit fly population doubles. On Day 6, her fruit fly habitat is full. On what day was it half full?"

Items were summed so that higher scores indicated analytic, deliberative thinking (1 = *Correct*, 0 = *Incorrect*).

Additional Measure Not Reported in Manuscript: Intolerance of Ambiguity

The *Intolerance of Ambiguity* subscale of *Need for Closure* (Roets & Van Hiel, 2007; Webster & Kruglanski, 1994) was administered in Studies 2, 3, and 4 as a potential moderator of the effects. However, this measure was not a significant predictor of social judgments or a significant moderator of social judgments, so it was not reported in the manuscript. The full results for these tests are reported below.

Results

Analyses Using Diffuse Priors

The same pattern of results emerged when using diffuse priors rather than informative priors based on the results of Study 1. Generic beliefs ($\beta = .451$ [.423, .478], $p < .001$) predicted social judgments more strongly ($b = -.116$ [-.166, -.065], $p < .001$) than statistical beliefs ($\beta = .335$ [.307, .362], $p < .001$).

Analyses Using Maximum-Likelihood Estimation

As in Study 1, we re-conducted the analyses in a two-level random-effects model (maximum likelihood, robust standard error estimation). Generic beliefs ($\beta = .470$, $SE = .024$, $t = 19.22$, $p < .001$) predicted social judgments more strongly ($b = -.117$, $SE = .004$, $t = -2.92$, $p = .004$) than statistical beliefs ($\beta = .352$, $SE = .024$, $t = 14.93$, $p < .001$). Results were no different to those conducted using Bayesian estimation.

Detailed Results of Moderating Effects

We tested the moderating effects of cognitive style and authoritarianism together *in the same model*, meaning that each moderator had to explain unique variance in social judgments (i.e., separate moderators could not be significant due to shared variance between cognitive style and authoritarianism).

Across studies, each of the other potential moderators (e.g., intolerance of ambiguity, participant age) was added by itself to this two-moderator base model and removed before entering the next moderator. Importantly, inclusion of these additional parameters did not alter the moderation results that we report in the manuscript.

Cognitive Style. Significant moderating effects emerged for cognitive style with statistical beliefs ($\beta = .035$ [.007, .064], $p = .008$) but not generic beliefs ($\beta = .005$ [-.024, .034], $p = .358$). Social judgments were more closely related to statistical beliefs for analytic thinkers (+1 *SD*; $b = .366$ [.324, .408], $p < .001$) than for intuitive thinkers (-1 *SD*; $b = .296$ [.257, .335], $p < .001$).

Authoritarianism. Significant moderating effects emerged for authoritarianism with both statistical ($\beta = -.048 [-.076, -.020], p = .001$) and generic beliefs ($\beta = .046 [.017, .075], p = .001$). Decomposing these interactions with tests of simple slopes, we found that (as predicted) high-authoritarian (+1 *SD*) participants' judgments ($b = .521 [.478, .565], p < .001$) relied on generic beliefs more than did low-authoritarian (-1 *SD*) participants' judgments ($b = .428 [.389, .467], p < .001$). In contrast (and also as predicted), high-authoritarian participants' judgments ($b = .283 [.242, .324], p < .001$) relied on statistical beliefs *less* than did low-authoritarian participants' judgments ($b = .379 [.339, .419], p < .001$).

Intolerance of Ambiguity. Intolerance of ambiguity did not directly predict social judgments ($\beta = .009 [-.024, .041], p = .29$). Intolerance of ambiguity also did not significantly moderate the extent to which statistical beliefs ($\beta = -.023 [-.050, .005], p = .055$) or generic beliefs ($\beta = .023 [-.005, .051], p = .052$) predicted social judgments. To remind the reader, Bayesian *p* values above .025 should be considered non-significant, particularly when models are appropriately powered (as is the case here), because Bayesian *p* values represent the proportion of the posterior distribution that overlaps zero, and therefore have a maximum value of .5.

Participant Age. Participant age was not correlated with authoritarianism ($r = .10, p = .19$) or cognitive style ($r = .09, p = .21$), the attitudinal and cognitive moderators that we expected would be prominent moderators of the links between beliefs and judgments. As in Study 1, participant age moderated the link between generic beliefs ($\beta = .037 [.008, .067], p < .001$), but not statistical beliefs ($\beta = -.008 [-.037, .021], p = .311$), and social judgments. The link between generic beliefs and judgments was stronger for older participants ($b = .500 [.456, .544], p < .001$) relative to younger participants ($b = .425 [.383, .466], p < .001$). This is again consistent with the idea that the low-effort nature of generic beliefs is part of the reason for their centrality in stereotypes.

Saliency of Stereotypes. Does the saliency of a stereotype (operationalized as frequency of mentions in our stereotype elicitation task) moderate the relationship of generic and statistical beliefs with social judgments? We assigned a *saliency score* to each stereotype according to how many participants listed that stereotype in the elicitation task. For example, “Black people are lazy” was the most commonly listed stereotype (14 mentions) and “Rock stars use drugs” was the least common (3 mentions). This saliency score (standardized) was a significant moderator. The more salient a stereotype, the stronger the relationship of generic beliefs ($\beta = .036$ [.009, .063], $p = .004$) and the weaker the relationship of statistical beliefs ($\beta = -.036$ [-.064, -.008], $p = .005$) with social judgments. In other words, generic beliefs were most influential for the most common, widely-known stereotypes. However, as we report below, the moderating effects of stereotype saliency did not replicate across studies.

STUDY 3

Method

Participants

Fifteen participants rating social items failed the attention check (excluded participants’ *average* prevalence estimate for “White people are humans” = 52.73%; included participants’ *average* > 99.7%). Ten participants rating non-social items failed the attention check (excluded participants’ *average* prevalence estimate for “Triangles have three sides” = 48.9%; included participants’ *average* > 99.9%). (We present percentages here for consistency. Recall, however, that participants did not provide percentages per se but rather rated “how many” members of a category displayed a trait on a sliding scale from *None* to *All*).

Selection of Items for Each Condition

Social stereotypes were identical to Study 2. We selected 34 non-social items from Khemlani et al. (2012). Items were selected from the categories of *striking*, *minority*

characteristic, majority false generalizations, and minority false generalizations (see Khemlani et al., 2012, for additional detail about these categories). We used 3 of these 34 items as counter-stereotypic fillers (e.g., “alligators have fur”) to match the composition of the social items and 1 item as an attention check (“triangles have three sides”). Of the 30 remaining non-social items, 15 were about animate entities (e.g., dogs, mosquitoes) and 15 were about inanimate entities (e.g., shirts, rooms).¹ Supplementary Tables 3 and 5 list the means and standard deviations for each of the individual social and non-social items, respectively. Supplementary Table 4 displays the aggregate means and standard deviations for the non-social items.

Results

Analyses Using Diffuse Priors

In the social condition, the same pattern of results emerged when using diffuse priors rather than informative priors based on the results of Study 2. Both statistical beliefs ($\beta = .332$ [.249, .371], $p < .001$) and generic beliefs ($\beta = .449$ [.410, .487], $p < .001$) predicted participants’ expectations about the traits of individual people. As in Studies 1 and 2, generic beliefs were a significantly stronger predictor of social judgments ($b = -.116$ [-.186, -.046], $p < .001$).

In the non-social condition, the analyses reported in the main text used diffuse priors because Study 3 was the only one in this paper to use non-social items.

Analyses Using Maximum-Likelihood Estimation

As in the prior studies, we re-conducted the analyses in a two-level random-effects model (maximum likelihood, robust standard error estimation). In the social condition, both

¹ Exploratory analyses indicated that the pattern of results reported in the main text was similar for animate and inanimate non-social items. That is, statistical beliefs were a strong predictor for both animate ($\beta = .627$ [.587, .671], $p < .001$) and inanimate ($\beta = .545$ [.493, .598], $p < .001$) items. Also similar to the main text, generic beliefs were a significant, but weaker, predictor for both animate ($\beta = .291$ [.250, .332], $p < .001$) and inanimate ($\beta = .290$ [.239, .342], $p < .001$) items.

generic beliefs ($\beta = .444, SE = .035, t = 12.69, p < .001$) and statistical beliefs ($\beta = .394, SE = .033, t = 12.06, p < .001$) predicted social judgments. These effects were not significantly different in magnitude ($b = -.049, SE = .060, t = -0.83, p = .41$), whereas a significant difference was detected when using Bayesian estimation that properly cross-classified participants and items.

In the non-social condition, statistical beliefs ($\beta = .590, SE = .032, t = 18.32, p < .001$) predicted social judgments more strongly ($b = .233, SE = .064, t = 3.67, p < .001$) than generic beliefs ($\beta = .356, SE = .034, t = 10.60, p < .001$). Results were no different from those conducted using Bayesian estimation.

Detailed Results of Moderating Effects

Cognitive Style. In the social condition, cognitive style did not predict judgments directly ($\beta = .019 [-.030, .068], p = .225$). Significant moderating effects emerged for cognitive style with generic beliefs ($\beta = -.057 [-.088, -.027], p < .001$) but not statistical beliefs ($\beta = .003 [-.029, .025], p = .423$). Analytic thinkers' expectations ($b = .401 [.359, .444], p < .001$) were less reliant on generic beliefs relative to intuitive thinkers' expectations ($b = .516 [.473, .559], p < .001$).

In the non-social condition, cognitive style similarly did not predict judgments directly ($\beta = -.005 [-.040, .029], p = .378$). A significant moderating effect emerged for cognitive style with generic beliefs ($\beta = -.057 [-.088, -.027], p < .001$): Analytic thinkers' expectations about non-social stimuli ($b = .341 [.303, .378], p < .001$) showed a weaker relationship with generic beliefs than did intuitive thinkers' expectations ($b = .448 [.406, .489], p < .001$). A significant moderating effect also emerged for cognitive style with statistical beliefs ($\beta = .003 [-.029, .025], p = .423$): Analytic thinkers' expectations about non-social stimuli ($b = .611 [.575, .646], p < .001$) were more strongly related to their statistical beliefs than were intuitive thinkers' expectations ($b = .461 [.421, .501], p < .001$).

This pattern was consistent with the patterns of moderation for cognitive style with social items across the other studies.

Authoritarianism. In the social condition, authoritarianism did not predict judgments directly ($\beta = -.011$, $[-.060, .039]$, $p = .327$). Significant moderating effects emerged for authoritarianism with generic beliefs ($\beta = -.074$ $[-.117, -.032]$, $p < .001$) but not statistical beliefs ($\beta = .036$ $[-.006, .078]$, $p = .043$). (As a reminder to the reader, Bayesian p values are considered significant under a threshold of .025.) The relationship between generic beliefs and expectations of social others was stronger for high-authoritarians ($b = .550$ $[.489, .612]$, $p < .001$) than for low-authoritarians ($b = .402$ $[.346, .457]$, $p < .001$).

In the non-social condition, authoritarianism did not predict judgments directly ($\beta = .009$, $[-.034, .053]$, $p = .338$), and did not moderate the relationship of either statistical beliefs ($\beta = -.024$ $[-.055, .008]$, $p = .067$) or generic beliefs ($\beta = .003$ $[-.029, .035]$, $p = .429$) with participants' non-social expectations. This pattern of effects was expected because authoritarianism is specifically related to people's orientation toward *social* groups—thinking about groups in all-or-none, hierarchical, and traditional ways.

Intolerance of Ambiguity. In the social condition, intolerance of ambiguity did not directly predict social judgments ($\beta = -.026$ $[-.078, .025]$, $p = .15$). Intolerance of ambiguity also did not significantly moderate the extent to which statistical beliefs ($\beta = -.026$ $[-.069, .017]$, $p = .113$) or generic beliefs ($\beta = -.024$ $[-.067, .019]$, $p = .136$) predicted social judgments.

Similarly, in the non-social condition, intolerance of ambiguity did not directly predict social judgments ($\beta = -.003$ $[-.039, .034]$, $p = .45$), nor did it moderate the extent to which statistical beliefs ($\beta = -.013$ $[-.046, .020]$, $p = .215$) and generic beliefs ($\beta = .018$ $[-.015, .052]$, $p = .144$) predicted non-social judgments.

Participant Age. As in Study 2, age was not correlated with authoritarianism ($r = .12$, $p = .23$) or cognitive style ($r = -.06$, $p = .53$). Also replicating the results of Study 2, in the social condition, participant age significantly moderated the relationship of generic beliefs ($\beta = .064$ [.026, .102], $p = .001$), but not statistical beliefs ($\beta = -.034$ [-.072, .003], $p = .036$), with social judgments. The link between generic beliefs and social judgments was heightened for older participants ($b = .519$ [.465, .572], $p < .001$) relative to younger participants ($b = .391$ [.333, .449], $p < .001$).

In the non-social condition, participant age did not moderate the relationship of generic beliefs ($\beta = -.002$ [-.032, .028], $p = .443$) or statistical beliefs ($\beta = -.017$ [-.046, .012], $p = .129$) with social judgments.

Salience of Stereotypes. In the social condition, the salience of the stereotypes (i.e., their frequency of mention in the elicitation task described in Study 2) did not predict judgments ($\beta = .011$ [-.030, .054], $p = .286$) or moderate the relationship of either generic beliefs ($\beta = .007$ [-.028, .043], $p = .346$) or statistical beliefs ($\beta = -.011$ [-.049, .025], $p = .278$) with social judgments.

This analysis was not conducted in the non-social condition because a measure of salience was not available for the non-social stimuli.

STUDY 4

Method

Participants

Twenty additional participants failed the attention check (*average* prevalence estimate for “White people are humans” = 42.7%; included participants’ *average* > 99.9%).

Results

Analyses Using Maximum-Likelihood Estimation

As in Study 1, we re-conducted the analyses in a two-level random-effects model (maximum likelihood, robust standard error estimation). Generic beliefs ($\beta = .142$, $SE = .023$, $t = 6.15$, $p < .001$) and statistical beliefs ($\beta = .123$, $SE = .020$, $t = 6.05$, $p < .001$) both predicted social judgments that the unfamiliar person's stereotypical trait had a biological basis. These effects were not significantly different in magnitude ($b = .018$, $SE = .039$, $t = 0.47$, $p = .64$). Results were similar to those conducted using Bayesian estimation, although did not detect a significant difference between the predictors, as was found when utilizing Bayesian estimation.

Detailed Results of Moderating Effects

Cognitive Style. Cognitive style moderated the relationship of generic beliefs ($\beta = -.054$ [$-.093, -.016$], $p = .003$) but not statistical beliefs ($\beta = .014$ [$-.024, .052$], $p = .232$) with participants' explanations. Tests of simple slopes revealed that biological explanations were more closely related to generic beliefs for intuitive thinkers ($b = .235$ [$.180, .290$], $p < .001$) than for analytic thinkers ($b = .127$ [$.071, .183$], $p < .001$).

Authoritarianism. A significant moderating effect emerged for authoritarianism with generic beliefs ($\beta = .074$ [$.037, .111$], $p < .001$) but not statistical beliefs ($\beta = .012$ [$-.026, .049$], $p = .269$). As expected, judgments that individuals' stereotypical traits were due to biology were more closely related to generic beliefs for people higher ($b = .255$ [$.201, .309$], $p < .001$) rather than lower ($b = .108$ [$.053, .161$], $p < .001$) in authoritarianism.

Intolerance of Ambiguity. Intolerance of ambiguity did not directly predict social judgments ($\beta < .0001$ [$-.012, .012$], $p = .50$). Intolerance of ambiguity also did not moderate the extent to which statistical beliefs ($\beta = -.003$ [$-.022, .015$], $p = .36$) or generic beliefs ($\beta = .007$ [$-.010, .024$], $p = .20$) predicted participants' biological explanations.

Participant Age. As in both prior studies, participant age was not correlated with authoritarianism ($r = .07$, $p = .37$) or cognitive style ($r = .03$, $p = .64$). However, unlike prior

studies, participant age showed no significant moderating effects. It did not moderate the links between statistical beliefs ($\beta = .019 [-.020, .059], p = .171$) or generic beliefs ($\beta = -.033 [-.073, .006], p = .050$) with participants' endorsement of biological explanations.

Summarizing the results across Studies 1–4, older (vs. younger) participants relied more heavily on *generic* beliefs in their social judgments when those judgments were expectations for individuals to have stereotypical traits (Studies 1–3), but not when those judgments were biological explanations for stereotypical traits (Study 4).

Salience of Stereotypes. The salience score of each stereotype moderated the relationship between statistical beliefs ($\beta = .082 [.040, .123], p < .001$), but not generic beliefs ($\beta = .040 [-.001, .080], p = .027$), and participants' biological explanations. However, the significant moderating effect for statistical beliefs was in the *opposite* direction of that identified in Study 2. Here, participants' biological explanations for stereotypical traits were *more* closely related to statistical beliefs for more (vs. less) salient stereotypes.

In summary, there were no robust or consistent moderating effects of stereotype salience across Studies 2–4. The relationship of *generic* beliefs with social judgments was stronger for more salient stereotypes in Study 2 but was not moderated by stereotype salience in Studies 3 or 4. The relationship of *statistical* beliefs with social judgments was weaker for more salient stereotypes in Study 2 but stronger in Study 4, and was not moderated by stereotype salience in Study 3.

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Supplementary Table 1
Means (and Standard Deviations) for the Items in Study 1

| Item | Statistical | Generic | Expectation |
|--|--------------------|----------------|--------------------|
| 1. African Americans are athletic | 56.63 (19.46) | 0.73 (1.05) | 0.60 (0.96) |
| 2. African Americans are musical | 51.00 (17.98) | 0.42 (1.05) | 0.33 (0.95) |
| 3. African Americans are undisciplined | 38.49 (22.59) | -0.60 (1.37) | -0.36 (1.36) |
| 4. African Americans are violent | 35.33 (22.59) | -0.58 (1.51) | -0.42 (1.35) |
| 5. Asians are ethnocentric | 52.29 (22.93) | 0.19 (1.13) | 0.22 (1.12) |
| 6. Asians are hardworking | 73.58 (14.96) | 1.33 (1.09) | 1.27 (1.03) |
| 7. Asians are intelligent | 70.22 (15.48) | 1.30 (1.10) | 1.14 (1.06) |
| 8. Asians are introverted | 51.25 (20.90) | 0.13 (1.13) | 0.28 (0.98) |
| 9. Caucasians are racist | 34.56 (21.23) | -0.79 (1.35) | -0.74 (1.33) |
| 10. Doctors are hardworking | 80.14 (14.13) | 1.75 (1.03) | 1.78 (1.00) |
| 11. Doctors are intelligent | 83.20 (14.86) | 1.91 (1.01) | 1.96 (0.98) |
| 12. Jewish people are cheap | 45.44 (24.44) | -0.12 (1.38) | 0.09 (1.29) |
| 13. Jewish people are ethnocentric | 50.22 (23.74) | 0.11 (1.27) | 0.11 (0.99) |
| 14. Jewish people are intelligent | 66.70 (16.94) | 0.96 (0.97) | 0.79 (0.88) |
| 15. Lawyers are competitive | 78.39 (15.79) | 1.68 (1.09) | 1.58 (0.98) |
| 16. Lawyers are dishonest | 50.25 (23.10) | 0.33 (1.39) | 0.28 (1.28) |
| 17. Lawyers are greedy | 58.19 (21.52) | 0.58 (1.35) | 0.59 (1.10) |
| 18. Men are assertive | 64.44 (16.06) | 0.96 (0.90) | 0.85 (0.85) |
| 19. Men are intellectual | 58.11 (16.97) | 0.70 (1.04) | 0.52 (0.89) |
| 20. Men are logical | 63.22 (15.22) | 0.80 (1.04) | 0.78 (0.96) |
| 21. Men are self-reliant | 65.28 (16.99) | 0.90 (0.92) | 0.98 (0.90) |
| 22. Politicians are dishonest | 64.26 (25.25) | 1.10 (1.37) | 0.92 (1.30) |
| 23. Teachers are hardworking | 74.87 (17.62) | 1.52 (1.10) | 1.41 (1.12) |
| 24. Teachers are intelligent | 73.94 (16.21) | 1.35 (1.10) | 1.35 (1.09) |
| 25. Women are compassionate | 69.97 (13.48) | 1.33 (1.03) | 1.16 (0.95) |
| 26. Women are honest | 62.95 (15.84) | 0.65 (1.07) | 0.47 (0.88) |
| 27. Women are sensitive | 68.35 (16.82) | 1.18 (1.07) | 1.12 (0.93) |

Note. The items are listed in generic form. *Statistical* = prevalence estimates (percentage). *Generic* = agreement with the generic statements. *Expectation* = social judgments of an individual (namely, the likelihood they will display a trait). Prevalence estimates were measured on 100-point sliding scales; other items were measured on -3 to 3 Likert-type scales.

Supplementary Table 2
Means (and Standard Deviations) for the Items in Study 2

| Item | Statistical | Generic | Expectation |
|-------------------------------------|--------------------|----------------|--------------------|
| 1. Asian people are bad drivers | 33.17 (25.14) | 34.35 (28.55) | -0.29 (1.44) |
| 2. Asian people are good at math | 56.65 (23.61) | 57.10 (25.83) | 0.81 (1.10) |
| 3. Asian people are smart | 62.05 (23.02) | 62.55 (24.34) | 0.82 (1.11) |
| 4. Black people are athletic | 45.36 (22.03) | 51.49 (23.96) | 0.28 (1.11) |
| 5. Black people are lazy | 26.39 (21.52) | 26.92 (25.80) | -0.82 (1.38) |
| 6. Black people are unintelligent | 23.80 (20.78) | 22.81 (25.43) | -0.99 (1.34) |
| 7. Black people are violent | 25.28 (21.38) | 29.91 (26.65) | -0.76 (1.36) |
| 8. Blondes are dumb | 25.24 (21.10) | 26.42 (25.34) | -0.84 (1.36) |
| 9. British people have bad teeth | 34.46 (23.50) | 37.55 (27.27) | -0.28 (1.42) |
| 10. Fat people are lazy | 41.69 (25.61) | 43.13 (28.92) | 0.03 (1.45) |
| 11. French people are snobs | 31.71 (22.96) | 36.89 (26.83) | -0.19 (1.40) |
| 12. Irish people are alcoholics | 27.38 (20.45) | 30.74 (26.26) | -0.55 (1.40) |
| 13. Jewish people are cheap | 34.71 (24.61) | 34.48 (26.64) | -0.34 (1.36) |
| 14. Jewish people are rich | 40.46 (23.08) | 43.70 (25.56) | -0.01 (1.25) |
| 15. Men are selfish | 43.19 (24.21) | 41.33 (25.29) | -0.18 (1.34) |
| 16. Men are strong | 57.68 (19.44) | 60.27 (21.49) | 0.70 (1.05) |
| 17. Mexicans are hard-working | 66.65 (20.38) | 63.86 (23.37) | 0.83 (1.10) |
| 18. Mexicans are lazy | 22.42 (18.04) | 22.10 (22.67) | -1.05 (1.23) |
| 19. Muslims are terrorists | 17.26 (21.91) | 26.46 (27.04) | -1.18 (1.54) |
| 20. Native Americans are alcoholics | 27.18 (21.77) | 30.18 (25.80) | -0.58 (1.44) |
| 21. Politicians are liars | 68.28 (25.86) | 71.84 (24.96) | 1.24 (1.40) |
| 22. Poor people are lazy | 28.93 (20.80) | 28.05 (25.63) | -0.66 (1.35) |
| 23. Redheads have a temper | 28.92 (22.7) | 30.37 (26.44) | -0.52 (1.29) |
| 24. Rock stars do drugs | 54.10 (23.72) | 58.15 (23.77) | 0.65 (1.19) |
| 25. Southerners are ignorant | 31.42 (23.71) | 31.66 (27.00) | -0.65 (1.39) |
| 26. Teenagers are irresponsible | 57.23 (25.01) | 61.90 (25.34) | 0.89 (1.27) |
| 27. White people are racist | 34.74 (22.85) | 38.08 (25.99) | -0.47 (1.37) |
| 28. White people are rich | 28.96 (20.05) | 35.12 (23.38) | -0.57 (1.16) |
| 29. Women are delicate | 41.22 (24.61) | 42.35 (27.88) | -0.06 (1.29) |
| 30. Women are unintelligent | 20.08 (16.04) | 16.72 (21.97) | -1.42 (1.29) |

Note. The items are listed in generic form. *Statistical* = prevalence estimates (percentage). *Generic* = agreement with the generic statements. *Expectation* = social judgments of an individual (namely, the likelihood they will display a trait). Prevalence estimations and agreement with the generic statements were measured on 100-point sliding scales; trait expectations were measured on -3 to 3 Likert-type scales.

Supplementary Table 3

Means (and Standard Deviations) for the Social Items in Study 3

| Item | Statistical | Generic | Expectation |
|-------------------------------------|--------------------|----------------|--------------------|
| 1. Asian people are bad drivers | 34.28 (22.87) | 37.06 (27.70) | -0.45 (1.42) |
| 2. Asian people are good at math | 55.89 (24.22) | 59.91 (22.95) | 0.63 (1.13) |
| 3. Asian people are smart | 59.87 (21.65) | 63.03 (20.57) | 0.63 (1.19) |
| 4. Black people are athletic | 47.13 (21.85) | 54.93 (24.00) | 0.35 (1.21) |
| 5. Black people are lazy | 29.83 (21.37) | 31.81 (26.77) | -0.72 (1.41) |
| 6. Black people are unintelligent | 27.90 (22.53) | 26.25 (25.97) | -1.03 (1.26) |
| 7. Black people are violent | 29.86 (22.85) | 35.06 (28.41) | -0.79 (1.33) |
| 8. Blondes are dumb | 26.14 (19.81) | 26.40 (23.89) | -0.92 (1.32) |
| 9. British people have bad teeth | 34.62 (22.49) | 42.58 (27.49) | -0.42 (1.32) |
| 10. Fat people are lazy | 43.64 (25.95) | 43.36 (27.53) | 0.11 (1.32) |
| 11. French people are snobs | 32.01 (23.83) | 41.71 (28.67) | -0.23 (1.37) |
| 12. Irish people are alcoholics | 30.01 (22.46) | 35.90 (26.07) | -0.56 (1.18) |
| 13. Jewish people are cheap | 38.89 (25.05) | 44.13 (29.37) | -0.20 (1.42) |
| 14. Jewish people are rich | 43.35 (25.26) | 47.55 (28.24) | -0.12 (1.39) |
| 15. Men are selfish | 42.77 (23.34) | 42.60 (25.10) | -0.38 (1.22) |
| 16. Men are strong | 55.77 (19.81) | 63.76 (19.47) | 0.66 (0.94) |
| 17. Mexicans are hard-working | 62.85 (20.92) | 64.95 (23.02) | 0.76 (1.17) |
| 18. Mexicans are lazy | 23.14 (19.52) | 21.16 (23.26) | -1.30 (1.21) |
| 19. Muslims are terrorists | 24.31 (27.66) | 29.66 (28.40) | -1.06 (1.63) |
| 20. Native Americans are alcoholics | 28.30 (21.75) | 30.56 (24.32) | -0.74 (1.30) |
| 21. Politicians are liars | 71.35 (26.34) | 75.21 (24.12) | 1.45 (1.37) |
| 22. Poor people are lazy | 34.40 (24.69) | 35.71 (28.33) | -0.52 (1.41) |
| 23. Redheads have a temper | 33.23 (24.30) | 38.82 (28.27) | -0.36 (1.49) |
| 24. Rock stars do drugs | 54.79 (25.17) | 61.77 (25.33) | 0.57 (1.31) |
| 25. Southerners are ignorant | 32.95 (22.94) | 31.20 (28.23) | -0.72 (1.46) |
| 26. Teenagers are irresponsible | 56.37 (24.66) | 60.99 (24.58) | 0.65 (1.28) |
| 27. White people are racist | 32.99 (21.21) | 35.36 (25.37) | -0.52 (1.23) |
| 28. White people are rich | 32.30 (21.09) | 34.82 (23.38) | -0.50 (1.26) |
| 29. Women are delicate | 46.46 (24.71) | 47.27 (25.95) | 0.12 (1.26) |
| 30. Women are unintelligent | 23.03 (21.05) | 15.92 (20.95) | -1.52 (1.23) |

Note. The items are listed in generic form. *Statistical* = prevalence estimates (percentage). *Generic* = agreement with the generic statements. *Expectation* = social judgments of an individual (namely, the likelihood they will display a trait). Prevalence estimations and agreement with the generic statements were measured on 100-point sliding scales; trait expectations were measured on -3 to 3 Likert-type scales.

Supplementary Table 4

Descriptive Statistics and Correlations across Measures for the Non-Social Items in Study 3

| | <i>Mean (SD)</i> | 1. | 2. | 3. | 4. |
|------------------------|----------------------------|------|------|-------|-------|
| 1. Statistical Beliefs | 49.59 (32.19) ^a | – | | | |
| 2. Generic Beliefs | 57.95 (33.06) ^a | .71* | – | | |
| 3. Expectation | 0.37 (1.83) ^b | .82* | .74* | – | |
| 4. Authoritarianism | -1.28 (1.08) ^b | .01 | .04 | .01 | – |
| 5. Cognitive Style | 1.59 (1.20) ^c | -.02 | .01 | -.04* | -.13* |

Note. Superscripts represent scale of measurement. *a* = 100-point sliding scales, *b* = -3 to 3 Likert-type scales, *c* = 0 to 3 scale. **p* < .05.

Supplementary Table 5
Means (and Standard Deviations) for the Non-Social Stimuli in Study 3

| Item | Statistical | Generic | Expectation |
|---|--------------------|----------------|--------------------|
| 1. Barns are red | 50.13 (24.14) | 58.83 (25.20) | 0.51 (1.31) |
| 2. Books are paperbacks | 56.76 (14.85) | 57.68 (21.12) | 0.58 (0.89) |
| 3. Cars have radios | 87.14 (11.93) | 87.26 (18.45) | 2.17 (0.95) |
| 4. Cats are white | 21.50 (13.66) | 38.44 (25.16) | -0.87 (1.22) |
| 5. Clocks are round | 62.62 (22.75) | 66.76 (20.82) | 1.05 (1.09) |
| 6. Computers are PCs | 66.79 (22.89) | 73.22 (23.46) | 1.32 (1.14) |
| 7. Dogs are blind | 11.76 (12.43) | 15.77 (18.71) | -1.97 (1.10) |
| 8. Dogs have tails | 91.21 (9.61) | 91.16 (15.16) | 2.45 (0.81) |
| 9. Ducks are female | 54.10 (9.04) | 51.70 (20.68) | 0.23 (0.71) |
| 10. Ducks lay eggs | 67.93 (24.25) | 80.55 (26.89) | 1.34 (1.64) |
| 11. Goats have horns | 62.88 (23.64) | 69.80 (26.44) | 1.07 (1.46) |
| 12. Hurricanes damage buildings | 64.46 (27.76) | 77.76 (24.40) | 1.52 (1.40) |
| 13. Jackets have zippers | 67.07 (19.71) | 73.77 (21.29) | 1.41 (1.07) |
| 14. Lions are male | 54.25 (16.64) | 55.89 (26.30) | 0.49 (1.13) |
| 15. Lions have manes | 60.54 (21.50) | 71.09 (26.44) | 1.04 (1.27) |
| 16. Mammals produce milk | 74.06 (25.55) | 80.60 (22.51) | 1.47 (1.42) |
| 17. Mosquitos carry malaria | 30.73 (22.51) | 60.31 (27.63) | -0.19 (1.61) |
| 18. Rats carry disease | 47.65 (27.43) | 67.81 (26.29) | 0.38 (1.50) |
| 19. Restaurants are Chinese restaurants | 23.71 (16.14) | 35.41 (25.40) | -0.52 (1.23) |
| 20. Rooms are round | 17.66 (17.66) | 23.84 (22.02) | -1.54 (1.26) |
| 21. Rottweilers maul children | 14.36 (14.71) | 29.68 (26.31) | -1.46 (1.34) |
| 22. Sharks attack swimmers | 22.16 (22.98) | 47.15 (30.88) | -0.57 (1.69) |
| 23. Sheep have udders | 43.97 (29.76) | 44.30 (31.95) | -0.02 (1.76) |
| 24. Shirts have collars | 56.63 (22.12) | 64.28 (22.00) | 0.86 (1.09) |
| 25. Shoes have laces | 64.48 (18.96) | 74.11 (22.32) | 1.21 (1.02) |
| 26. Tables are 10 feet long | 24.27 (18.06) | 37.46 (25.61) | -0.87 (1.24) |
| 27. Ticks carry Lyme disease | 39.28 (25.67) | 64.19 (29.07) | 0.12 (1.62) |
| 28. Tigers are albino | 12.01 (15.58) | 26.53 (26.80) | -1.79 (1.23) |
| 29. Trees are deciduous | 49.54 (27.34) | 55.19 (29.23) | 0.57 (1.29) |
| 30. Trumpets are loud | 84.85 (18.29) | 84.67 (20.16) | 2.09 (1.10) |

Note. The items were adapted from Khemlani et al. (2012) and are listed in generic form. *Statistical* = prevalence estimates (estimate of how many of the category members have that property). *Generic* = agreement with the generic statements. *Expectation* = predictions about an individual (namely, the likelihood they will display a trait). Prevalence estimations and agreement with the generic statements were measured on 100-point sliding scales; trait expectations were measured on -3 to 3 Likert-type scales.

Supplementary Table 6
Means (and Standard Deviations) for the Items in Study 4

| Item | Statistical | Generic | Biol Expl |
|-------------------------------------|--------------------|----------------|------------------|
| 1. Asian people are bad drivers | 28.15 (20.42) | 28.54 (26.61) | -2.33 (1.11) |
| 2. Asian people are good at math | 52.20 (22.04) | 54.19 (27.16) | -0.94 (1.85) |
| 3. Asian people are smart | 55.20 (22.31) | 58.06 (26.23) | -0.60 (1.87) |
| 4. Black people are athletic | 42.37 (21.02) | 49.48 (25.56) | 0.29 (1.79) |
| 5. Black people are lazy | 23.45 (20.27) | 20.60 (24.42) | -2.12 (1.46) |
| 6. Black people are unintelligent | 20.37 (19.25) | 16.51 (23.42) | -1.52 (1.75) |
| 7. Black people are violent | 23.75 (19.78) | 25.53 (27.60) | -1.56 (1.70) |
| 8. Blondes are dumb | 19.98 (17.39) | 18.27 (21.79) | -1.53 (1.70) |
| 9. British people have bad teeth | 31.31 (19.90) | 32.01 (25.77) | -0.33 (2.01) |
| 10. Fat people are lazy | 37.02 (25.07) | 35.97 (28.53) | -1.69 (1.53) |
| 11. French people are snobs | 28.49 (21.73) | 29.93 (25.79) | -2.31 (1.13) |
| 12. Irish people are alcoholics | 24.85 (19.23) | 26.40 (24.76) | -0.99 (1.68) |
| 13. Jewish people are cheap | 27.54 (22.15) | 26.59 (26.02) | -2.19 (1.33) |
| 14. Jewish people are rich | 33.40 (22.02) | 36.75 (26.78) | -2.38 (1.10) |
| 15. Men are selfish | 39.13 (21.80) | 37.91 (25.97) | -1.91 (1.34) |
| 16. Men are strong | 55.21 (20.36) | 58.70 (23.86) | 0.70 (1.67) |
| 17. Mexicans are hard-working | 64.96 (18.14) | 63.71 (23.09) | -1.84 (1.52) |
| 18. Mexicans are lazy | 18.47 (15.36) | 16.39 (19.47) | -2.19 (1.27) |
| 19. Muslims are terrorists | 17.35 (21.70) | 24.48 (26.31) | -2.33 (1.29) |
| 20. Native Americans are alcoholics | 24.81 (19.08) | 24.15 (23.81) | -0.99 (1.80) |
| 21. Politicians are liars | 67.56 (26.23) | 69.08 (26.83) | -2.18 (1.25) |
| 22. Poor people are lazy | 27.48 (21.66) | 23.89 (25.25) | -2.14 (1.30) |
| 23. Redheads have a temper | 27.33 (21.54) | 26.43 (25.50) | -1.20 (1.75) |
| 24. Rock stars do drugs | 51.34 (23.79) | 57.25 (25.71) | -1.72 (1.64) |
| 25. Southerners are ignorant | 28.01 (21.14) | 28.96 (26.68) | -2.17 (1.22) |
| 26. Teenagers are irresponsible | 56.54 (21.84) | 59.47 (24.70) | -1.03 (1.82) |
| 27. White people are racist | 31.43 (22.22) | 32.74 (26.21) | -2.34 (1.26) |
| 28. White people are rich | 24.83 (17.35) | 30.51 (24.22) | -2.31 (1.26) |
| 29. Women are delicate | 36.67 (22.15) | 36.64 (28.09) | -0.29 (1.83) |
| 30. Women are unintelligent | 18.78 (16.83) | 10.48 (17.34) | -1.45 (1.81) |

Note. The items are listed in generic form. *Statistical* = prevalence estimates (percentage). *Generic* = agreement with the generic statements. *Biol Expl* = social judgments of an individual (namely, endorsement of a biological explanation for the stereotypical trait). Prevalence estimations and agreement with the generic statements were measured on 100-point sliding scales; endorsement of biological explanations was measured on -3 to 3 Likert-type scales.